

Discrimination in Mortgage Lending: Evidence from a Correspondence Experiment

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Abstract:

We design and implement an experimental test for differential response by Mortgage Loan Originators (MLOs) to requests for information about loans. Our e-mail correspondence experiment is designed to analyze differential treatment by client race and credit score. Our results show net discrimination by 1.8 percent of MLOs through non-response. We also find that MLOs offer more details about loans and are more likely to send follow up correspondence to whites. The effect of being African American on MLO response is equivalent to the effect of having a credit score that is 71 points lower.

Keywords: Discrimination, Field Experiment, Mortgage Lending, Race

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I. Introduction

Allegations of discriminatory lending practices during the 2004-2008 housing boom have resulted in the two largest cash settlements ever between mortgage lenders and the Department of Justice (DOJ). The settlements, \$335 million from Bank of America's Countrywide group and \$175 million from Wells Fargo,¹ allege that these institutions steered equally qualified minority applicants into higher interest (sub-prime) loans and charged higher fees than for white borrowers. During this same time period, Home Mortgage Disclosure Act (HMDA) data released by the Federal Financial Institutions Examination Council² shows only an 8 basis point difference in the average interest rate charged to white and non-white borrowers (favoring whites).³ The substantial cost of discriminatory lawsuits and the lack of corroboration in the aggregate data begs the question, "do mortgage lenders *really* discriminate against minority borrowers?" and if so, "how is this possible in an age of computerized, nearly automatic, underwriting?"

We address these questions by testing for racial discrimination using a matched-pair correspondence experiment on Mortgage Loan Originators (MLOs). MLOs are essentially licensed mortgage sales workers who assist customers with loan applications and have the ability to offer and negotiate the terms of a mortgage with applicants. The role of information provider and advisor in the lending process, and the discretion MLOs have in dealing with customers makes them an integral part of the borrowing process from a client's perspective. Discrimination by MLOs could result in different lending outcomes between minority and majority borrowers,

¹ The Countrywide and Wells-Fargo lawsuits were settled in December, 2011 and July, 2012, respectively. The DOJ alleges that Countrywide charged higher fees and rates to more than 200,000 African American and Hispanic borrowers and steered more than 10,000 African American and Hispanic borrowers into high interest mortgages (New York Times, December 21st, 2011). Similarly, the DOJ alleges that Wells-Fargo charged higher fees and rates to approximately 30,000 African American and Hispanic borrowers, and steered more than 4,000 African American and Hispanic borrowers into high interest mortgages (DOJ, July 12th, 2012). It is typical that lending institutions do not admit guilt when settling a charge of discrimination with the DOJ.

² The FFIEC maintains summary statistics of HMDA data on its website at: <http://www.ffiec.gov/hmdaadwebreport/NatAggWelcome.aspx>.

³ The 8 basis point difference between white and non-white borrowers is not conditional on borrower characteristics and is the difference in the mean interest rate reported for loans where the interest rate is known on conventional, 1-4 family home purchase loans (excluding manufactured homes) between 2004 and 2008. Whites have an approximately 27 basis point differential (favoring whites) with African American borrowers over this time period, not conditional on borrower characteristics. See Gruenstein-Bocian et al. (2008) for a study that examines interest rate differences across race groups conditional on borrower characteristics.

and also influence outcomes as the home buying process proceeds. For example, a borrower who is delayed or who is pre-approved for a smaller loan amount may be treated differently by a real estate agent in terms of search effort, neighborhood choice, or expediency of service. If differences in initial treatment by an MLO are severe (offering different interest rates, fees, or suggesting credit repair services), this could conceivably affect a home buyer in all aspects of the home purchase, even if they are successful in obtaining a loan.⁴

Our matched-pair experiment examines the response MLOs offer to initial contact from a potential client interested in obtaining information about a mortgage loan. We design the experiment to test for differential treatment by client race (white or African American) and by credit score. We randomly assign pairs of e-mail inquiries to MLOs according to our design to test for the effects of a borrower's race, credit score, and the interaction between these two. We reveal client race to MLOs using selected client names within each e-mail inquiry. We use only names that have a high likelihood of being given to only one race in a sample of birth certificate for male babies born in New York City in 1990. We examine both the propensity for MLOs to respond to our inquiries, the propensity to follow up, and the content of the response to test for differential treatment.

To our knowledge, this is the first experimental test of discrimination by MLOs that uses e-mail correspondence.⁵ This is in contrast to earlier studies by Smith and Delair (1999) and Ross et al. (2008), that rely on in-person interaction between MLOs and actors. Heckman (1998) and Heckman and Siegelman (1993) critique using actors when testing for discrimination because actors may bias results if they are not identical along all dimensions except race. The severity of actor bias can be diminished by carefully choosing and training actors, but actors may also be subject to misreporting or inadvertently prompting a discriminatory response (Hanson and Hawley, 2011).

⁴ See Ross and Yinger (2002) for a particularly lucid explanation of discrimination in the lending process, including an explanation and critique of research methodology.

⁵ There are several recent studies that use e-mail correspondence to test for discrimination in the market for rental housing. See Hanson and Hawley (2011) for a recent example and a review of this literature. Also see Ladd (1998), Yinger and Ross (2002), and Ross et al. (2008) for a review of the literature on discrimination in mortgage markets in particular

While we believe there is value in using in-person studies, and they offer ways to examine discrimination by MLOs that our study cannot, our work provides some advantages over in-person studies.⁶ Most importantly, we avoid actor bias by relying solely on electronic communication with MLOs, allowing us to dramatically increase the scope of the experiment and the geographic area covered relative to in-person studies. Using electronic communication also provides a detailed record of correspondence which allows us to examine the timing and content of MLO responses to our inquiries.

The use of the internet in general is becoming a standard part of the home search and borrowing process which has yet to receive much attention in the academic literature. Bricker et al. (2010) report that 41.7 percent of borrowers use the internet for information about borrowing,⁷ and over 90 percent of home buyers in 2012 reported using the internet in some capacity during their home search (NAR, 2012).

Our results show that MLOs discriminate on the basis of race and treat clients differently by their reported credit score. We find that on net, 1.8 percent of MLOs discriminate by not responding to inquiries from African Americans while responding to inquiries from white clients.⁸ We find larger net response differences across credit score types, with 8.5 percent of MLOs responding to clients in our high credit score group while not responding to clients who do not report a credit score. We also find that credit score differences exacerbate differences in response between races. Overall, the effect of being African American on MLO response is roughly equivalent to the effect of having a credit score that is 71 points lower. Examining the content of the response shows that whites are favored even among MLOs that respond to both inquiries. The primary difference in the content of response between whites and African

⁶ See Doleac and Stein (2013) for a novel approach to avoiding the use of actors by studying discrimination using pictures in an on-line market. This work varies the skin color of the seller to test for discrimination among buyers of iPods.

⁷ In addition to the internet, 39.5 percent of borrowers report using sellers of financial services as a method of obtaining information about borrowing. The most commonly used source of information about borrowing is “friends, relatives, and associates” with 43.9 percent of borrowers using that channel (Bricker et al., 2010).

⁸ The net level of discrimination measures the difference in the percentage of MLOs that only reply to an inquiry from a white client against the percentage of MLOs that only reply to an inquiry from an African American client. The gross level of discrimination or the percentage of MLOs that only reply to an inquiry from a white client is 17.8 percent of MLOs. The overall difference in response rates is 2.6 percentage points favoring whites- this difference does not match the net discrimination level because some of our experiments involved sending inquiries from same race clients to the same MLO.

Americans is the inclusion of details about the loan. We also find that MLOs are more likely to send follow-up correspondence to whites than African Americans.

The remainder of the paper is organized as follows: Section II describes an MLO's role in the lending process. Section III outlines the design of our experiment, while section IV details implementation and sample characteristics. Section V presents our primary results while Section VI offers several robustness checks. The final section of the paper offers concluding comments.

II. MLOs in the Lending Process

MLOs are typically the initial and primary contact person for borrowers seeking a mortgage, and have discretion over how they respond to customer inquiries. MLOs may, for example suggest that a borrower attempt to improve their credit score before completing a loan application, or may encourage a borrower to act quickly to take advantage of low interest rates. They may also present different fees or interest rates to borrowers, offer encouragement or discourage the borrower from moving forward with the loan, or offer other financial advice related to obtaining a mortgage.⁹

MLOs typically have contact with the client throughout the entire lending process, from initial inquiry through loan closing, but they are particularly important in the application process. Clients who have marginal credit depend on MLOs to give advice on what products to apply for, what steps to take to improve their credit and whether their application will ultimately be successful. MLOs may communicate with an underwriter, but do not directly make decisions about accepting or denying a loan. Ross et al. (2008) point out that while minorities are less likely than whites (controlling for observable factors) to obtain a loan,¹⁰ it is quite rare to find discrimination in mortgage underwriting. Ross et al. suggest that discrimination by MLOs would be consistent with these facts.

⁹ The Secure and Fair Enforcement for Mortgage Licensing Act (SAFE), part of the larger Housing and Economic Recovery Act of 2008, included several provisions to tighten regulations of MLOs. These provisions included requiring licensing of MLOs, creating a Nationwide Mortgage Licensing System (NMLS), issuing uniform licensing applications and reporting requirements across states, and creating a national clearing house for collecting consumer complaints.

¹⁰ See Munnell et al. (1996) for a study that identifies denial rate differences between African American and white clients controlling for credit differences.

III. Experiment Design

To test for discrimination among MLOs we design a matched pair correspondence experiment using e-mail to inquire about assistance with a home mortgage.¹¹ The matched pairs are structured to test outcome differences due to race and credit score differences among potential borrowers. Each MLO receives two e-mails in the experiment. This design, along with the use of names to identify race, follows the Bertrand and Mullainathan (2004) resume experiment. The audit methodology has a long history in the housing discrimination literature starting with Yinger's (1986) real estate agent experiment.

We use three credit score groups in our experiment: no credit score, low credit score, and high credit score. The low credit score group reports a randomly assigned credit score between 600 and 650; the high credit score group reports a randomly assigned credit score between 700 and 750.¹² As a precaution against exposing the experiment, we randomly assign a credit score for each e-mail (rather than each pair) from a uniform distribution within each category (low or high). Although there is a chance the credit scores within a matched pair are exactly the same, most often the scores will be different within a small range. We use the randomly assigned differences in credit scores to test how MLOs respond to credit score and race differences. For the portion of experiments where a credit score is reported, the average credit score is 675. For the high score group the average is 725 for the low score group it is 625.

We also divide our experiment into groups by the content of correspondence. We chose this design to guard against exposing the experiment. Each group includes one of two types of questions to be asked of the MLO. Using different questions across MLOs may make our inquiries less suspicious for company spam filters or for co-workers who discuss client e-mails. E-mails to one group contain a question about interest rates and a question about mortgage fees (all questions for this set are listed in Appendix 1 in the boxes labeled Question #1a and

¹¹ The experiment attempts to uncover discrimination from differential treatment based on minority status. Scholars (and governments) have also recognized that disparate impact, or having a policy that disproportionately impacts minorities while lacking business purpose is discrimination. See Turner and Skidmore (1999) for a discussion about the difference between differential treatment and disparate impact in mortgage lending.

¹² The low credit score range approximates the 15-30th percentile of the national distribution of credit scores according to the Fair Isaac Company (FICO). The high credit score range approximates the 40-60th percentile of the national FICO score distribution (FRB, 2007). Most MLOs seem to operate using a rule on an acceptable credit score (like a minimum of 620); we noticed that the reported rule varied across the responses we received. Our low score sample seems to straddle the rule in all areas.

Question #1b). E-mails to the other group contain a question about loan availability and a question about what information is necessary to proceed in the process of obtaining a loan (all questions for this set are listed in Appendix 1 in the boxes labeled Question #2a and Question #2b). To further guard against exposing the experiment to MLOs, we randomly assign the phrases within the structure of our e-mail inquiries. For example, we randomly assign each e-mail one of five possible greetings (Hello, Hi, Hi There, Hey, or Dear), and ensure that the other e-mail sent to the same MLO does not use the exact greeting. We view the benefits from not matching the text exactly (reducing the risk of exposing the experiment) as exceeding the cost that any of our greetings (or other text elements) might influence outcomes in a meaningful way.¹³ Appendix 1 details the exact layout of our correspondence and the randomly assigned text that populates each e-mail.

Our experiment includes 30 different matched pair types, representing all of the combinations between cells in Figure 1.¹⁴ This allows us to examine the marginal effect of race and credit score (on the extensive and intensive margin), as well as to examine if there is a different marginal effect of credit score across races. We randomly assign each MLO to a matched pair type, and randomly vary the credit score within the range for that type. The matched pair, or within-subjects, design means that each MLO in our experiment receives two e-mail inquiries.

We reveal borrower race to MLOs through the name associated with each e-mail inquiry. The source of first names is the New York City Department of Health and Human Hygiene (DHHH) records for babies born in 1990. The DHHH birth records provide counts of babies born by gender, race, and first name. We begin by calculating the probability a baby is born either white or African American for each name in the sample. We use only male baby names for this calculation. The DHHH data do not report a count for names with fewer than 10 babies born in a given race-gender match. This makes our probabilities for names that are very likely to be associated with only one race equal to one, when in fact they could be less than one. Because of

¹³ There is almost no difference in the response rate across types of e-mails. The response rate for the question set #1 group is 67.2. The response rate for the question set #2 group is 66.8. This difference is not statistically meaningful.

¹⁴ We also randomly vary the order in which e-mails are sent. For each matched pair type the order of the treatment difference (e.g. high vs. low credit score) is randomly reversed in exactly half the emails. We do this to ensure that that order effects do not drive any results.

this censoring, and the primary concern of signaling race, we also consider the raw number of occurrences each name has within a given race. After compiling a list based on probabilities and counts, we eliminate most names that have a Muslim or Jewish origin from our list as we want to minimize any confounding effects these characteristics would bring to the experiment.

The source of surnames is Word et al.'s analysis of 2000 Census data. This analysis reports counts of surnames for the general population, and by race/ethnicity of respondents to the census. For African American surnames we use the same criteria as first names, choosing those with the highest probability of belonging to African Americans. We choose the surnames with the largest probability of belonging to African Americans regardless of total count, as the data shows a large number of African Americans with these surnames in all cases. We use slightly different criteria for white surnames, as many of the names with the highest probabilities of belonging to whites have a strong ethnic component (for example the highest probabilities are Yoder, Mueller, Koch, all are from a German origin). For white surnames, we choose three names (Miller, Nelson, Baker) from the most common (by count) names that have greater than a 0.8 probability of being white and less than a .15 probability of being African American. We choose the other two names (Krueger and Schmitt) from the list with the highest probabilities of being white, regardless of their ethnic attachment.

Table 1 shows the list of names used to signal race in the experiment. The first three columns of Table 1 show the probability a baby is African American or white given they are born with that name, the count of babies born with that name in 1990, and the rank (by count) for each name. The last three columns of Table 1 show the probability a person is African American or white given the surname, the count of persons with that name in 2000, and the rank (by count) for each name. White names chosen with the alternative criteria are not ranked in the top ten for all white names, thus we report their value for rank as NA. MLOs are exposed to the name associated with each inquiry in three ways: the actual e-mail address,¹⁵ the signature at the bottom of each e-mail (styled as "First name Surname"), and the name plate in the MLOs inbox (styled as "First name Surname).

¹⁵ We use Gmail addresses exclusively, and all take the form `firstname.surnameXXX@gmail.com`, where XXX is a random three-digit number.

Table 2 shows the frequency each name occurs in the experiment. Each name represents approximately 5 percent of the sample of e-mails sent, or about 520 e-mails. The least frequent name in our sample is Kadeem Jefferson, with 459 e-mails or about 4.4 percent of the total e-mails in our experiment. The most frequent name in our sample is Tyrone Washington, with 577 e-mails or about 5.6 percent of the total e-mails in our experiment. All differences in name frequency are due to the random assignment of names to matched-pair types, and random assignment of matched-pair types to MLOs in our experiment.

IV. Experiment Implementation and Sample Characteristics

We identify a set of MLOs with their e-mail addresses as subjects for the experiment through internet search. Collection of MLO contact information occurred from February through April of 2012. We used multiple styles of internet search including Google Maps, Google.com, Yellow Pages (YP.com) and Better Business Bureau (bbb.org). For each MLO we identify the following information: name (first and last), state, website, email address, title, physical address and company affiliation. When a photograph is available, we also identify their presumed gender and race.

To ensure a broad sample and limit the potential for the experiment to be exposed, we limit sampling of MLOs operating in the same workplace. We categorize MLO workplace according to their place of employment on two levels: the company and the branch. We consider MLOs to work for the same company if they work for an employer with the same company name (for example, Bank of America or Wells Fargo). We consider MLOs to work for the same branch if they advertise the same physical address on their website. We limit our sample to 8 MLOs per branch, but do not restrict the number of MLOs at the company level. The average number of MLOs per branch in our experiment is 2.38, with an average per company of 27.

We balance the number of MLOs by state level geography, using the proportion of the US population in 2010. For instance, Mississippi has 0.96 percent of the 2010 US population and we target 0.96 percent (50) of MLOs in our sample to come from that state. Our goal is not perfect geographic representativeness, but rather a broad geographic sample of the target

population of MLOs. Table 3 shows a state-by-state count of the MLOs in our experiment.¹⁶ The state with the largest number of subjects is California with 423 MLOs, whereas Alaska and West Virginia have the smallest number, 2 each. The difference between the proportion of MLOs in our sample and 2010 sample population ranges from under sampling by 3.9 percentage points in California to over sampling by 1.71 percentage points in Illinois. Most states are within 0.50 percentage points of the population proportion.

We have 5,181 as subjects in our experiment. A notable feature of MLOs is their relative demographic homogeneity. We are able to identify the race of MLO using photographs for about 75 percent of our sample. 93 percent of the race-identified MLOs are white. Gender is more evenly split, where 53 percent of gender-identified MLOs are male. We know of no existing demographic statistics to corroborate our sample as representative of the industry. Nonetheless, we did not specifically seek MLOs based on race or gender, and believe these statistics to be representative of the industry participants who list information on the internet. Table 4 shows complete demographic characteristics for the sample of MLOs in the experiment.

Our experiment began on Monday, April 30, 2012 at 2:00 pm with a set of 200 pilot emails to a sub-sample of MLOs. The full experiment commenced one week later, on Tuesday, May 8 at 1:00 pm. An MLO was randomly selected to participate in one of five rounds (one pilot and four regular).¹⁷ A round consists of a first and second e-mail going out to the same group of MLOs- separated by one week. For instance, Round 3 recipients received their first email on Thursday, May 10; they received their second email on Thursday, May 17. All emails following the pilot were sent beginning at 1pm Eastern Daylight Time (EDT) and ending by 2pm EDT. Regular rounds were conducted on Tuesdays, Wednesdays and Thursdays between May 8 and May 29. The schedule is designed to minimize possible confounding effects. Mondays and Fridays are avoided in the experiment to minimize the impact of weekend lag times or end-of-

¹⁶ The actual sample in our experiment is not exactly proportional with 2010 state populations. This is partly due to randomly selecting MLOs from this sample to be subjects and partly due to the fact that availability of MLO information is not uniform across geography.

¹⁷ Random selection for a round depends on the number of MLOs at a branch. Spreading a branch's MLOs through multiple rounds is done to ensure no more than two emails from any of our clients arrive at an office on any given day. The structure of this selection, however, is independent of treatment assignment.

week effects. The start time of each half round is held constant to minimize unobserved time-of-day effects.

Keeping days of the week and time of e-mail sending similar within rounds is designed to reduce noise in the experiment. Bias is eliminated through randomization of treatment assignment and treatment order. Whichever audit type an MLO is assigned to, the order in which the MLO receives the treatment is randomized. Should our efforts to eliminate confounding effects be inadequate, randomizing the order of treatment causes such effects to attenuate the outcome to zero rather than bias it. This is a key strength of our experimental design.

Given the volume of mail to be delivered (two emails to each MLO from one of twenty different client email accounts), we used an automated email-sending program to minimize human error and speed processing. The program was set to initiate individual Simple Mail Transfer Protocol (SMTP) sessions for each email sent with the Gmail servers, and to wait one to two seconds between sending emails. The resulting emails are indistinguishable from messages that would be sent directly from the Gmail web interface.

Of the 10,362 e-mail inquiries sent in our experiment, about 1.3 percent received an automated response or out-of-office reply. We do not count these as a response unless a follow up e-mail was sent by the MLO.

V. Results

The experiment provides a wealth of information on MLOs response to standard inquiries for assistance with obtaining a home mortgage. We maintain all content for each MLO response to use in the analysis, in addition to the date/time stamp that the e-mail was sent to examine speed of response.

Response vs. Non-Response by Race

The most basic indication of equal treatment across race and credit score types is whether or not an MLO responds to our inquiry for assistance with a mortgage loan. We consider a response to be one that we deem genuinely written by a human; out-of-office or other automatic replies are not considered a response. Our measure of response is if the MLO ever sent a

genuine e-mail response within two weeks of our inquiry. This measure means that we count a genuinely human response to have occurred even if it is not the first correspondence we received from the MLO.¹⁸ In several instances, we received out of office replies, and subsequent genuinely human responses. In other cases we received automatically generated commercial replies followed by genuinely human responses.

Table 5 shows the difference in genuine response by MLOs to our inquiries across race groups. The first and second columns of Table 5 show the overall response rate, while columns 4-7 offer a breakdown at the MLO level for the sub-sample of the experiment that includes MLOs who received separate e-mails from both an African American and a white borrower.¹⁹ The response rate to white borrowers for the full sample is 68.31 percent. For African American borrowers, the response rate is 2.63 percentage points lower, or 65.68 percent. As shown in the third column of Table 5, this difference is statistically significant at the one percent level, indicating that MLOs are less likely to respond to inquiries from African American borrowers than they are from whites.

We find a slightly higher response rate difference (3.26 percentage points in favor of white borrowers) for the sub sample of MLOs whose race is white. The magnitude of our results is similar for the sub-sample of non-white MLOs, but the small sample size of this group strains statistical precision. We find a smaller difference in response rate for the sample of MLOs where we are not able to identify race (a 0.9 percentage point difference, favoring whites), although this result is also imprecise. The level of discrimination measured by response rate differences between male and female MLOs is similar, 3.16 percentage points for male MLOs, 3.57 percentage points for females, both favoring white borrowers and both precisely estimated. Among the sample of MLOs where we do not identify gender, we find only a small, imprecisely estimated difference in response rate (0.19 percentage points, favoring African Americans).

Although we do find a statistically significant difference in response rates between African American and white borrowers, the MLO level results shown in columns 4-7 of Table 5 show

¹⁸ Of the responses we consider genuinely human, 99.8 percent of them were received within two weeks of our original inquiry. We count the other 0.02 percent as non-responses.

¹⁹ Recall that the full sample includes some audits where inquiries had the same race, but different credit characteristics.

that 66.05 percent of the MLOs in our sample treat e-mail inquiries the same- either by responding to both (49.77) or responding to neither (16.28) inquiry. We measure discrimination at the MLO level by the net amount of discrimination, or the difference in the proportion of MLOs who respond only to whites, and those that only respond to African Americans. Using this measure, we find a smaller level of discrimination: MLOs responding only to whites outnumber MLOs responding only to African Americans by 1.81 percentage points, a difference which is close to statistically precise at conventional levels (a p-value of 0.0573) despite the smaller sample size and more rigorous significance test.²⁰

The MLO-level results across race and gender of MLOs are similar to the response rate results, with smaller magnitudes and some loss of statistical precision. We still find that white MLOs discriminate, but only by 2.14 percentage points on net, with statistical significance at the 10 percent level. Our results for male MLOs also remain largely the same, showing a slightly higher level of discrimination than the full sample (2.89 percentage points on net), and maintaining statistical precision; however, we lose statistical precision on the results for female MLOs.

Response vs. Non-Response by Credit Score

Table 6 shows the difference in genuine response by MLOs to our inquiries across credit score groups. Panel A shows the overall response rate difference across the high, low, and no credit score groups. Panel B shows the difference in response at the MLO level for the sub-sample of MLOs who received separate e-mails from borrowers with different credit scores (or one including a credit score and the other excluding). The high credit score group received the highest response rate, 69.46 percent. The response rate for the low credit score group is 3.7 percentage points lower, or 65.76 percent. The difference in response rate between the high and low credit score group is statistically significant at the one percent level, indicating that MLOs are more likely to respond to inquiries from borrowers with a higher credit score.

²⁰ For all statistical significance tests at the MLO level, we use the McNemar test. This test is designed for testing the difference in proportion of respondents for paired subjects. The test statistic is $\chi^2 = (N_{\text{only } W} - N_{\text{only } AA})^2 / (N_{\text{only } W} + N_{\text{only } AA})$, where N represents the number of MLOs only responding to one group. The McNemar test statistic has a chi-squared distribution, and we calculate all p-values accordingly.

We find a similar gap in response rates between the high and no credit score groups of 3.69 percentage points, with the response rate of the no credit score group at 65.77 percent. The difference in response rate between the high and no credit score group is also statistically significant at the one percent level. We find only a small, statistically imprecise gap between the response rate for the low and no credit score groups that slightly favors the no credit score group.

The MLO level results show that the majority of MLOs respond (between 49.88 and 53 percent) or do not respond (between 15.17 and 17.8 percent) to both credit score groups. As with the race results, we measure equal treatment by examining the net proportion of MLOs that respond differently across credit score groups. The MLO level analysis reveals about the same level of differential treatment between the high and low credit score groups: 4.03 percentage points on net, which is statistically significant at the 10 percent level (p value of 0.0506). The biggest difference is the differential treatment between the high and no credit score group, where the high credit score group is favored by 8.48 percent of MLOs on net. This result is statistically meaningful at the one percent level.

Comparing the race and credit score differences shows that MLOs are relatively more sensitive to differences in credit scores when deciding whether or not to respond to a borrower inquiry for assistance with a mortgage loan. The relative difference between race and credit score groups depends on which response measure is used. At the MLO level, going from a low to high credit score roughly doubles the difference in net unequal treatment in response/non-response compared to the difference between African American and white borrowers. The mean difference between credit score groups (100 points), assuming a linear relationship between credit score and response, suggests that the effect of having an African American name on MLO response is roughly equivalent to having a credit score that is 71 points lower.

Response vs. Non-Response by Race and Credit Score

The design of our experiment allows us to look at differential treatment across several possible race-credit score combinations. Table 7 shows results for audits where whites and African Americans are assigned to the same credit score category, where whites are in a higher credit score category, and where African Americans are in a higher credit score category. The magnitude of the response rate differences and MLO level differences will be exactly the same

for these results, as they use the same set of audits; however, the statistical tests at the MLO level count only differential response by individual MLOs, so statistical significance may vary.

The magnitude of discrimination using only audits in equal credit score categories is slightly smaller than the level for the entire sample; we find net discrimination by 1.62 percent of MLOs, as opposed to 1.9 for the full sample. The statistical significance of this relationship is strained, as the MLO level result has a p-value of 0.1693, outside of traditional significance levels, while the response rate test is close to statistical precision at the 10 percent level. Given that the magnitude of these results is similar to the full sample, the loss of statistical precision does not seem alarming, especially considering that these results rely on a smaller sample.²¹

The results across race where clients report different credit scores show that the higher credit score group is favored regardless of whether they are white or African American, but that the degree to which they are favored is larger for whites with higher credit scores. The middle row of Table 7 shows that whites are favored by 10.21 percent (statistically significant at the one percent level) of MLOs in audits where the white client reports a higher credit score and the African American reports either a low or no credit score. The bottom row of Table 7 shows that while African Americans with higher credit scores are favored over whites with low or no credit score, the difference is only 7.11 percent of MLOs on net (statistically significant at the 5 percent level). The MLO level results show that more MLOs choose to reply to both clients when the African American has a higher credit score, as opposed to replying to only the white client when African Americans have a lower credit score. These results suggest that while there is a level of discrimination that exists regardless of credit score, discrimination increases with credit score differences between races. A non-constant level of discrimination may mean that the source of discrimination is not taste-based, but instead a form of statistical discrimination based on perceived group differences.

Content of Response

E-mail communication with MLOs allows us to examine not only the propensity to respond, but also the nature of responses in our experiment. Table 8 shows how MLOs

²¹ We also test for differences when both scores are high, both low, and both do not include credit. These results show a similar level of discrimination as the results that combine these groups, but statistical precision is even more strained due to smaller sample sizes in these sub groups.

responded in terms of the timing, length, and propensity to send a follow-up response in the experiment.

Results show that MLOs are substantially slower to respond to African American clients than they are to whites. Among MLOs that responded to both inquiries, whites received a response in 8 hours and 20 minutes on average. MLOs took an average of an hour and three minutes longer to respond to African Americans, a result that is statistically precise at the ten percent level. In terms of the length of response, a measure of response intensity, we find no difference in the number of characters in a response when examining all e-mails. Examining a sub-set of responses that differ by at least 10 characters (to eliminate short generic, or form responses), we do find that whites are favored and receive a response that is about 4 percent longer.

The propensity to attempt a follow up with white potential clients was much greater than for African American clients. We find that about 7 percent of MLOs sent at least a second reply to whites, while only 5.2 percent sent a second reply to African Americans. Statistically, this is one of our strongest results, as it is statistically meaningful at the one percent level. We find, however, that the average number of replies sent to the two groups is not different.

Content of Response: Side by Side Comparisons

We designed a side-by-side analysis tool that allows us to make a direct comparison of the response from a single MLO to clients with a different race. Appendix 2 shows a screen shot of the side-by-side analysis tool used to grade the difference between two e-mails sent by the same MLO to different clients. To ensure an unbiased grading, all identifying information from both MLOs and clients was masked when using the side-by-side analysis tool.

We conducted both an internal (author examined) and external²² (a team that did not include any authors) reviews of e-mail pairs using this tool. We designed the analysis tool to randomly assign a left-side and right-side e-mail, to guard against any ordering effects in grading responses. Graders were instructed to indicate if they thought the responses were “Neutral”, meaning they were similar in content, language, and nature; they felt one was “Preferred” in

²² We used a team of 10 different reviewers.

some way over the other, or they thought that one was “Strongly Preferred” over the other. Graders were instructed to strictly record their opinion about the response, and that they did not have to justify their feelings. In addition to an opinion about how favorable the responses were relative to one another, we also offered a series of check boxes for the reason. Graders were instructed that they were not required to use a reason, and there was also a place to write in other reasons. Reasons for favoring included: offering more favorable terms, friendliness, and facilitation of the transaction. Appendix 3 provides the instructions given to graders and the full list of reasons for preference.

We graded all responses from MLOs replying to both inquires for mortgage assistance. For white/African American matched pairs, this was 1,932 pairs of responses or 3,864 e-mails. Panel A of Table 9 shows the results of the side-by-side analysis for author graded responses. We find that even among MLOs that respond to both inquiries for mortgage assistance, some discrimination exists. Blind grading shows that about 57 percent of e-mails were perceived as being neutral between the different inquiries. Blind grading also shows that among e-mails where some preference was indicated, we were more likely to perceive e-mails sent to whites (22.6 percent) as being favorable to those sent to African Americans (20 percent). The magnitude of this difference suggests that another 2.6 percent of MLOs discriminates by responding more favorably to whites- this is in addition to the 1.8 percent of MLOs that discriminate through non-response. The side-by-side comparison results are statistically precise at the 10 percent level with a p-value of 0.0817.

The most common reason for choosing an e-mail was preferred for white clients was that they were given more details (46.6 percent of white preferred e-mails), this reason was also the most common for preference given to African Americans, but occurred only 40.8 percent of the time. The second most common reason for choosing a white client was preferred was that the tone of the e-mail was more friendly (33.56 percent), which was also the second most common reason African Americans were preferred (32.8 percent). It was also fairly common to choose that preference was given to whites because the e-mail facilitated the loan transaction (27.8 percent), although among MLOs that gave preference to African Americans this occurred more often (30.7 percent). It was less common for graders to indicate that preference was given to

whites because of strong overt measures of discrimination like offering more favorable terms (5.9 percent)²³ or steering into a product or being pushy (1.4 percent).

The side-by-side comparison done by outside reviewers is remarkably consistent with our internal grading. Although outside reviewers graded e-mails as neutral less often than the authors (43.9 percent, as opposed to 57.2), the net level of discrimination for favoring whites is only 0.15 percentage points higher. Panel B of Table 8 shows tests for discrimination in the content of response using the outside reviewer's opinion of MLO replies. The outside reviewers perceived that e-mails to whites were more favorable 29.2 percent of the time, while perceiving favorable e-mails for African Americans 26.4 percent of the time. The net incidence of unequal treatment is 2.8 percent of MLOs, but this result is not quite statistically significant with a p-value of 0.106. Outside reviewers were more likely to use friendliness and the inclusion of details as reasons why whites were favored and less likely to use the "other" category.

Why do MLOs Discriminate?

Although our experiment is not specifically designed to test for the reasons why MLOs discriminate against African American clients, the data generated from our experiment allows us to explore this question to some degree. The standard theories behind why discrimination occurs fall into two basic categories: taste-based (Becker, 1957), and statistical (Phelps, 1972). Ewens et al. (in press) outline a model and several hypotheses that allow distinction between the two types of discrimination in an experimental setting in the market for rental housing. The basic premise in Ewens et al. is that positive or negative information will be viewed differently by agents possessing limited information, resulting in differences in treatment between groups across information types. Their formal hypothesis is that if agents practice taste based discrimination, then including positive information with an inquiry will amplify the racial gap that exists when no such information is given. According to their model, including negative information should narrow the gap, regardless of the type of discrimination practiced.

We examine the data from our experiment using the Ewens et al. framework, by testing for differences in treatment of African Americans with high and low credit scores relative to a

²³ Graders were asked to indicate that more favorable terms were offered if the lender replied with more favorable loan terms to one recipient than the other. This could be in a quoted or suggested interest rate, length of loan, type of loan, fees, or anything else that has to do with costs to the borrower.

baseline with no credit score reported. Our results show that low credit score African Americans are treated worse and high credit score African Americans are treated better than the omitted credit score group, which is suggestive that MLOs are not practicing taste-based discrimination in our experiment.

Bertrand et al. (2005) suggests an alternative to the standard taste-based and statistical explanations for discrimination. They suggest that agents may not make a conscious choice when discriminating, but instead discriminate unintentionally or implicitly because of an unconscious association between a person of a certain type and some identified attribute. This type of discrimination is tested in the laboratory by showing subjects photographs of people in a rapid manner and requiring that they assign them to some category (e.g. good or bad). Requiring quick reaction attempts to identify the subconscious thoughts of the subject.

Our experiment does not lend itself to a formal test of implicit discrimination, but the nature of the timing in MLO responses is suggestive. Table 8 shows a large gap between the average time it takes an MLO to respond to our subjects- they respond faster to whites. This is suggestive that on the margin, more MLOs are making the quick decision to respond to whites and to set aside the e-mail from the African American until later, which suggests the motivation for discrimination may be implicit rather than conscious discrimination.

VI. Robustness of Findings

Identifying discrimination in correspondence experiments relies on the choice of names being representative of race groups. The birth certificate data we use demonstrates that the names in our experiment are highly correlated with either race, but they leave open the possibility that particular names are treated differently for other reasons. While we do not have the ability to infer why this might be, we can examine how our choice of names may affect our results.

We consider two sets of robustness checks with the names in the experiment. First, we examine response rate differences between each name and other names of the same race, and exclude names that are treated statistically different. Next we examine national data on name popularity and exclude the most and least popular names in each race.

Table 10 shows the response rate by name in the experiment. Since our findings show that African Americans experience lower response rates, most of the top response rates come from white names. We perform t-tests of the response rate differences between each name and the average within-race response rate; these results are shown in the second column of Table 10. There are only two statistically distinguishable (among own race) names in our experiment: Jake Krueger has a higher response rate than other whites ($p=0.0422$) and Maxwell Baker has a lower response rate than other whites ($p=0.0901$).

Excluding white names with a higher than average response rate (Jake Krueger) narrows the gap between white and African American response rates, as shown in column 1 of Table 11. We are somewhat confident in these results as both tests show a similar (albeit smaller) gap between whites and African Americans, and the response rate tests maintain statistical significance. The within MLO results do not maintain statistical significance, but this is likely due to sample size restriction that comes with excluding 10 percent of the already smaller sample. Results excluding the other name that showed statistically different outcomes, Maxwell Baker, are in line with our primary results, and in fact show a slightly larger response rate difference. The third row shows that if we exclude white name with the highest, and African American name with the lowest response rate the magnitude of our results is smaller but still statistically meaningful for the response rate difference, but loses marginal statistical significance for the MLO level test.

We also examine differences in general popularity of the names in our experiment using Social Security Administrative data. This data reports counts of name for babies born nationally each year (unconditional on race) for the 1,000 most common names. All names in our experiment are in among the 1,000 most common in popularity for the year of our birth certificate data (1990). As a result of population shares, white names are necessarily more popular among the general population, but one of the white names (Conor) is quite unpopular nationally relative to other white names. The third column of Table 10 shows popularity ranks for all names in our data.

Excluding the most popular white name (Zackary Miller) and the least popular African American name (DaShawn Banks) does not affect either the response rate or MLO level results, as shown in Table 11. In each case, the magnitude of the difference is extremely similar and we maintain the same level of significance as our primary results. We also explore excluding the

two most popular white names, and (separately) excluding the two least popular African American names. In either case the response rate results are similar in magnitude to our primary results and maintain statistical significance. The MLO level results also show a similar magnitude to the primary results, but lose statistical significance.

Our overall interpretation of the robustness checks is that our results are fairly robust to excluding particular names, especially when the choice is based on popularity of the names and not directly a function of response.

VII. Conclusion

We find evidence of discrimination against African Americans in the market for mortgage loans. The discrimination we find occurs at the initial information gathering stage for borrowers in response to a simple e-mail inquiry about assistance with obtaining a mortgage. We find that MLOs, the primary contact person for a borrower looking to obtain a mortgage, are less likely to respond to inquiries from clients with African American names than they are to clients with white names. We also find that MLOs responding to inquiries from both races are more likely to write a preferential e-mail to white clients. The level of discrimination we find is large for a characteristic that should not matter (race) relative to one that should matter (credit score).

Finding discrimination in the information gathering stage is likely to influence outcomes for minority borrowers throughout the lending and home buying process. If African American borrowers are less likely to receive communication from an MLO and the MLO treats them differently when communication does occur, it makes submitting a loan application more difficult, and the remainder of the home purchase more arduous. In addition, our work shows that the growing importance of e-mail communication between clients and lenders, where in-person meetings are less and less common, does not mean that discrimination on the basis of race will not occur.

The magnitude of discrimination we find is smaller than the most recent in-person study (Ross et al., 2008); however, the standard for compliance is much lower in our most basic test: we only examine if MLOs are willing to respond to an e-mail. Our findings confirm that discrimination still exists in the lending industry, and that it exists across a larger sample, and geographic scope than previous studies have examined. We are also able to compare the

difference in treatment between whites and African Americans with the difference in treatment across credit score groups. Our average differences suggest an African American name reduces the probability that an MLO responds by the same magnitude as does reporting a credit score that is 71 points lower.

Our results suggest that examining lending outcomes is not sufficient to uncover the level of discrimination that minorities face in the lending process. Our work also suggests that to uncover the full extent of discrimination in this market, multiple types of communication should be used in addition to in-person audits, and that enforcement of Fair Lending Laws would be more robust if audits included other means of communication.

References

- Becker, Gary (1957) *The Economics of Discrimination*, Chicago: University of Chicago Press.
- Bertrand, Marianne, Mullainathan, Sendhil (2004) "Are Emily and Greg more employable than Lakisha and Jamal? A field experiment on labor market discrimination." *American Economic Review* 94 (4), 991–1013.
- Bricker, Jesse, Arthur Kennickell, Kevin Moore, and John Sabelhaus (2010) "Changes in U.S. Family Finances from 2007 to 2010: Evidence from the Survey of Consumer Finances" *Federal Reserve Bulletin* 98(2).
- Department of Justice, Office of Public Affairs Press Release. "Justice Department Reaches Settlement with Wells Fargo Resulting in More Than \$175 Million in Relief for Homeowners to Resolve Fair Lending Claims." July 12th, 2012.
- Doleac, Jennifer and Luke Stein (2013) "The Visible Hand: Race and Online Market Outcomes" *Economic Journal*, 123: 469–492.
- Ewens, Michael, Bryan Tomlin, and Choon Wang "Statistical Discrimination or Prejudice? A Large Sample Field Experiment" Forthcoming in the *Review of Economics and Statistics*.
- Federal Reserve Board (2007) "Report to the Congress on Credit Scoring and its effects on the Availability and Affordability of Credit".
- Gruenstein Bocian, Debbie, Keith Ernst and Wei Li (2008) "Race, ethnicity and subprime home loan pricing," *Journal of Urban Economics*, 60(1): 110-124
- Hanson, Andrew and Zackary Hawley (2011) "Do landlords discriminate in the rental housing market? Evidence from an internet field experiment in US cities," *Journal of Urban Economics*, 70 (2-3): 99-114.
- Hanson, Andrew, Zackary Hawley and Aryn Taylor (2011) "Subtle Discrimination in the Rental Housing Market: Evidence from E-mail Correspondence with Landlords," *Journal of Housing Economics* 20(4): 276 –284.
- Heckman, James (1998) "Detecting Discrimination." *Journal of Economic Perspectives*, 12(2): 101–16.
- Heckman, James and Peter Siegelman (1993) "The Urban Institute Audit Studies: Their Methods and Findings." In *Clear and convincing evidence: Measurement of discrimination in America*, ed. Michael E. Fix and Raymond J. Struyk, 187–258. Washington: Urban Institute Press.
- Ladd, Helen (1998) "Evidence on Discrimination in Mortgage Lending," *Journal of Economic Perspectives* 12(2): 41-62.

Munnell, Alicia, Geoffrey Tootell, Lynn Browne, and James McEneaney (1996) "Mortgage Lending in Boston: Interpreting HMDA Data," *American Economic Review*, 86 (1): 25-53.

National Association of Realtors, "Profile of Home Buyers and Sellers, 2012."

O'Toole, James. "Wells Fargo in \$175M discriminatory lending settlement" Money.CNN.com, July 12th, 2012.

Phelps, Edmund (1972) "The Statistical Theory of Racism and Sexism," *American Economic Review* 62(4), 659-661.

Ross, Stephen L., Margery Austin Turner, Erin Godfrey, Robin R. Smith (2008) "Mortgage lending in Chicago and Los Angeles: A paired testing study of the pre-application process," *Journal of Urban Economics*, 63 (3): 902-919.

Ross, Stephen L., and John Yinger (2002) "The Color of Credit: Mortgage Discrimination, Research Methodology, and Fair-Lending Enforcement" Cambridge: MIT Press.

Savage, Charlie. "Countrywide Will Settle a Bias Suit" *New York Times*, December 21st, 2011.

Smith, Robin and Michelle Delair (1999) "New Evidence from Lender Testing: Discrimination at the Pre-Application Stage. In: Turner, Margery and Felicity Skidmore (Eds.), *Mortgage Lending Discrimination: A Review of Existing Evidence*. Washington, D.C. Urban Institute Press.

Turner, Margery and Felicity Skidmore (1999) "Introduction, Summary, and Recommendations" In: Turner, Margery and Felicity Skidmore (Eds.), *Mortgage Lending Discrimination: A Review of Existing Evidence*. Washington, D.C. Urban Institute Press.

Word, David L., Charles D. Coleman, Robert Nunziata, and Robert Kominski.(not dated)"Demographic Aspects of Surnames from Census 2000. Technical Report for the U.S. Census Bureau

Yinger, John, (1986) Measuring racial discrimination with fair housing audits: caught in the act. *American Economic Review*, 76 (5), 881-893.

Figure 1: Experiment Design

		Questionset #1		Questionset #2	
		Black	White	Black	White
No CS	1	2	7	8	
Low CS	3	4	9	10	
High CS	5	6	11	12	

Appendix 1: Correspondence Construction

[GREETING] [FIRST NAME]²⁴,
 I'm interested in [PRODUCT].
 [SOURCE]
 [PLEASANTRY]
 [*If score known, then* CREDIT SCORE] [RANDOMLY ASSIGNED SCORE]
 [QUESTION #1a or QUESTION #2a, depending on question set type]
 [QUESTION #1b or QUESTION #2b, depending on question set type]
 [VALEDICTION]
 [RANDOMLY ASSIGNED NAME from Table X]

<p><u>GREETING</u> Hello Hi Hi There Hey Dear</p>	<p><u>PRODUCT</u> a home loan. a mortgage. getting a home loan. getting a loan. information on mortgages.</p>	<p><u>SOURCE</u> I found your information on-line, and thought you could help. I got your contact information on-line, and hope you can help me. I looked you up on-line, hopefully you can help. I found you on-line, and think you can help. I got your information on the web, and thought you might be able to help.</p>
<p><u>PLEASANTRY</u> I just have a few questions. I have a few questions for you. I'm curious about a few things. I'd like to ask you a couple of questions. I'm wondering about a few things.</p>	<p><u>CREDIT SCORE</u> I know that my credit score is My credit score is I have a credit score of I already know my credit score is</p>	<p><u>VALEDICTION</u> Thank you for your time, Thanks in advance, I look forward to hearing from you, I look forward to your reply, Thanks for taking the time,</p>
<p><u>QUESTION #1a</u> How are interest rates looking? What interest rate can I expect? Can you tell me about current interest rates? What do interest rates look like right now? How should I expect interest rates to look?</p>	<p><u>QUESTION #2a</u> What types of loans might be available for me? Can you tell me about the types of loans you have? What sort of loans are available for someone like me? Can you offer advice on what type of loans are available? What kind of loans do you have available?</p>	
<p><u>QUESTION #1b</u> What sort of fees are involved? What fees should I expect? How do the fees work? Are there fees that I need to worry about? What are the typical fees?</p>	<p><u>QUESTION #2b</u> What other information do you need from me? Do you need any other information to start the process? What sort of information do you need to move forward? What more do you need from me to proceed? Do you know what else I need to begin the process?</p>	

²⁴ We use actual MLO first names given on the webpage where we found contact information.

Appendix 2: Side-by-Side Comparison Tool

Comparison Form

<p>Hi XXXXX..</p> <p>Thanks for emailing me.. What I would need is for you to go to my website and fill out an application . This will give me permission to pull credit.</p> <p>Let me know that you see my website when I send you this email. I'm sending it from my iPhone.</p> <p>Once I pull credit we can talk about different loans. But the score you gave will give you many options. How much did you want to put down? FHA is 3.5 % down .. Let me know.</p> <p>If my website is not attached I can send it.</p>	<p>Hello XXXXX,</p> <p>I could help you with a loan that would fit best for you. We're you thinking to put as little as 3.5% or more down... Once I know this and your credit score we can discuss what would be best for you.</p> <p>Let me know if you are ready to have your credit pulled and get pre approved for a loan.</p> <p>Thanks!</p> <p>AAAAA</p>								
<p>Please indicate which email you prefer:</p> <p><input type="radio"/> Strongly Prefer <input type="radio"/> Prefer <input type="radio"/> Neutral <input type="radio"/> Disprefer <input type="radio"/> Strongly Disprefer</p> <p>Please enter your ID: <input type="text" value="ARH"/></p>									
<p>Please indicate the reasons for your preference. Check all that apply:</p> <table border="0"><tr><td><input type="checkbox"/> More favorable terms (interest rate, etc.)</td><td><input type="checkbox"/> Unrelated email was more negative</td></tr><tr><td><input checked="" type="checkbox"/> Friendliness</td><td><input type="checkbox"/> Facilitated the transaction</td></tr><tr><td><input type="checkbox"/> Included more details</td><td><input type="checkbox"/> Unrelated email steered into a product or was pushy</td></tr><tr><td><input type="checkbox"/> Explained the process</td><td><input type="checkbox"/> Other: <input type="text"/></td></tr></table> <p><input type="button" value="Previous"/> <input type="button" value="Next"/></p>		<input type="checkbox"/> More favorable terms (interest rate, etc.)	<input type="checkbox"/> Unrelated email was more negative	<input checked="" type="checkbox"/> Friendliness	<input type="checkbox"/> Facilitated the transaction	<input type="checkbox"/> Included more details	<input type="checkbox"/> Unrelated email steered into a product or was pushy	<input type="checkbox"/> Explained the process	<input type="checkbox"/> Other: <input type="text"/>
<input type="checkbox"/> More favorable terms (interest rate, etc.)	<input type="checkbox"/> Unrelated email was more negative								
<input checked="" type="checkbox"/> Friendliness	<input type="checkbox"/> Facilitated the transaction								
<input type="checkbox"/> Included more details	<input type="checkbox"/> Unrelated email steered into a product or was pushy								
<input type="checkbox"/> Explained the process	<input type="checkbox"/> Other: <input type="text"/>								

Appendix 3: Grader Instructions for Side-by-Side Comparison

Thank you for agreeing to help out with this research project. Your task is simple- to review a set of e-mail responses we received from mortgage lenders and compare them. We are interested to know if you feel the responses you read tend to be more favorable toward one set of recipients than another, or if they are treated approximately the same. Essentially, we want to know your opinion. Please use the format we've supplied by enabling external content in excel and simply clicking on the "Open Form" button. After entering the form, please start by typing your name in the ID box. Next, read both e-mails **carefully**. After reading both e-mails, please indicate if you thought the mortgage lender strongly preferred one, preferred the other, or treated the recipients neutrally using the check box indicator.

If you thought that the mortgage lender expressed a preference, please use the next section of check boxes to indicate why you felt this way. Feel free to use the "other" check box in the event that your opinion does not match a reason listed, or if you can't quite describe why you feel that way. You can also fill in a reason for "other" to describe your reasoning. Please use the following as guidance when checking boxes for your reasons:

More favorable terms (interest rate, etc.): Check this box if the lender replied with more favorable loan terms to one recipient than the other. This could be in a quoted or suggested interest rate, length of loan, type of loan, fees, or anything else that has to do with costs to the borrower.

Friendliness: Check this box if you feel the lender was more 'friendly' to one recipient than the other. Again, this is your opinion, we will not hold it against you.

Included more details: Check this box if you feel the lender gave a more detailed description of the products, application materials needed, or generally gave answers with more depth to one recipient than the other.

Explained the process: Check this box if you feel the lender offered more guidance on the lending process, the application process, or the home purchase process to one recipient than the other. This might include offers on how to improve credit, or necessary paper work to complete an application.

Un-preferred email was more negative: Check this box if you thought that one of the e-mails was negative, even if the other e-mail was neutral. This might include negative language, unusually short replies (relative to the other), or a rude tone in writing.

Facilitated the transaction: Check this box if you feel the lender attempted to facilitate a successful transaction more with one recipient than the other. This might include offers for future communication, providing application materials, encouraging an application, or offering help with credit issues or home search.

Un-preferred email steered into a product or was pushy: Check this box if you feel the lender was being pushy about selling a loan, or suggested a specific product that was "right" for the recipient and not the other. Differentiating between this box and facilitating a transaction will largely depend on your interpretation of the language of the e-mail. Remember, this is your opinion.

Table 1. Names Identifying Race

	First Name			Last Name		
	P(Race Name)	Count	Rank	P(Race Name)	Count	Rank
<u>White</u>						
Zachary Miller	1	164	1	0.86	969910	NA
Brendan Nelson	1	55	5	0.8	329788	NA
Jake Krueger	1	43	9	0.97	36694	2
Ethan Schmitt	1	38	10	0.97	35326	6
Maxwell Baker	1	36	15	0.82	343081	NA
Spencer Miller	1	31	17	0.86	969910	NA
Brett Nelson	1	28	20	0.8	329788	NA
Conor Schmitt	1	21	33 (tie)	0.97	35326	6
Luke Krueger	1	22	31	0.97	36694	2
Seth Baker	1	21	33 (tie)	0.82	343081	NA
<u>African American</u>						
Jamal Washington	1	96	1	0.9	163036	1
Jerome Jefferson	1	38	27	0.53	666125	2
DaQuan Booker	1	68	10	0.66	35101	3
Terrell Banks	1	66	12	0.54	99294	4
Darnell Jackson	1	65	13	0.53	666125	5
Tyrone Washington	1	56	14	0.9	163036	1
Kadeem Jefferson	1	84	2	0.75	51361	2
Reginald Jackson	1	51	18	0.75	51361	5
Jermaine Booker	1	49	22	0.66	35101	3
DaShawn Banks	1	39	26	0.54	99294	4

Notes: The source of first names is the New York City Department of Health and Human Hygiene (DHHH) records for babies born in 1990. The DHHH data do not report a count for names with fewer than 10 babies born in a given race-gender match. This makes our probabilities for names that are very likely to be associated with only one race equal to one, when in fact they could be less than one. The first name count is the number of babies born with that name for each race. The first name rank is where each name ranks in the count distribution.

The source of surnames is Word et al.'s analysis of 2000 Census data. For African Americans we choose the surnames with the largest probability of belonging to African Americans regardless of total count. The white surnames Krueger and Schmitt were chosen with the same criteria. Because we are concerned that many of the highest probability white surnames have a German origin, we choose three white surnames (Miller, Nelson, and Baker) using alternative criteria. The alternative criteria is to use the most common (by count) names that have greater than a 0.8 probability of being white and less than a .15 probability of being African American. White names chosen with the alternative criteria are not ranked in the top ten for all white names, thus we report their value for rank as NA.

Table 2. Frequency of Names in Experiment

	Frequency of occurrence	Percentage of e-mails
<u>White</u>		
Zachary Miller	509	4.91%
Brendan Nelson	535	5.16%
Jake Krueger	526	5.08%
Ethan Schmitt	528	5.10%
Maxwell Baker	515	4.97%
Spencer Miller	489	4.72%
Brett Nelson	502	4.84%
Conor Schmitt	508	4.90%
Luke Krueger	547	5.28%
Seth Baker	523	5.05%
<u>African American</u>		
Jamal Washington	513	4.95%
Jerome Jefferson	571	5.51%
DaQuan Booker	543	5.24%
Terrell Banks	518	5.00%
Darnell Jackson	485	4.68%
Tyrone Washington	577	5.57%
Kadeem Jefferson	459	4.43%
Reginald Jackson	490	4.73%
Jermaine Booker	497	4.80%
DaShawn Banks	527	5.09%
Total	10362	

Notes: Names are randomly assigned to an audit type after the audit type is randomly determined for each MLO. Random assignment is done without replacement for each MLO so that names are not repeated within an audit. Difference in frequency of names in the experiment is due to random assignment.

Table 3. Number of Audits and Response Rate across States

	Number of audits	Overall response rate (%)	Responded to at least one inquiry (%)	Percent of Audit Population	Percent of US Population
Full Sample	5181	68.50%	84.93%	--	--
Alabama	70	70.71%	88.57%	1.35%	1.55%
Alaska	2	75.00%	100.00%	0.04%	0.23%
Arizona	125	64.00%	81.60%	2.41%	2.07%
Arkansas	49	57.14%	89.80%	0.95%	0.94%
California	423	69.27%	83.69%	8.16%	12.07%
Colorado	99	70.20%	86.87%	1.91%	1.63%
Connecticut	69	57.97%	73.91%	1.33%	1.16%
Delaware	15	63.33%	80.00%	0.29%	0.29%
Florida	226	73.01%	86.28%	4.36%	6.09%
Georgia	224	71.43%	86.61%	4.32%	3.14%
Hawaii	28	60.71%	71.43%	0.54%	0.44%
Idaho	34	69.12%	88.24%	0.66%	0.51%
Illinois	304	70.89%	83.88%	5.87%	4.16%
Indiana	143	75.87%	92.31%	2.76%	2.10%
Iowa	60	69.17%	85.00%	1.16%	0.99%
Kansas	37	74.32%	89.19%	0.71%	0.92%
Kentucky	93	70.97%	83.87%	1.80%	1.41%
Louisiana	70	72.14%	84.29%	1.35%	1.47%
Maine	44	80.68%	95.45%	0.85%	0.43%
Maryland	158	70.25%	85.44%	3.05%	1.87%
Massachusetts	106	51.89%	68.87%	2.05%	2.12%
Michigan	231	73.81%	91.77%	4.46%	3.20%
Minnesota	95	72.63%	86.32%	1.83%	1.72%
Mississippi	57	71.05%	85.96%	1.10%	0.96%
Missouri	101	65.35%	87.13%	1.95%	1.94%
Montana	16	87.50%	100.00%	0.31%	0.32%
Nebraska	51	82.35%	94.12%	0.98%	0.59%
Nevada	57	73.68%	94.74%	1.10%	0.87%
New Hampshire	54	68.52%	79.63%	1.04%	0.43%
New Jersey	111	73.87%	82.88%	2.14%	2.85%
New Mexico	48	65.63%	83.33%	0.93%	0.67%
New York	328	68.75%	83.23%	6.33%	6.28%
North Carolina	64	74.22%	90.63%	1.24%	3.09%
North Dakota	27	61.11%	74.07%	0.52%	0.22%
Ohio	261	72.22%	91.19%	5.04%	3.74%
Oklahoma	64	47.66%	87.50%	1.24%	1.22%
Oregon	100	68.50%	83.00%	1.93%	1.24%
Pennsylvania	177	77.97%	89.27%	3.42%	4.11%
Rhode Island	34	73.53%	82.35%	0.66%	0.34%
South Carolina	103	63.59%	83.50%	1.99%	1.50%
South Dakota	12	45.83%	66.67%	0.23%	0.26%
Tennessee	58	68.97%	87.93%	1.12%	2.06%
Texas	225	52.44%	80.44%	4.34%	8.14%
Utah	85	46.47%	62.35%	1.64%	0.90%
Vermont	15	83.33%	100.00%	0.29%	0.20%
Virginia	144	58.33%	75.69%	2.78%	2.59%
Washington	159	68.87%	86.16%	3.07%	2.18%
West Virginia	2	75.00%	100.00%	0.04%	0.60%
Wisconsin	110	76.82%	90.91%	2.12%	1.84%
Wyoming	13	50.00%	76.92%	0.25%	0.18%

Notes: Sampling is intended to follow U.S. state population proportions. The exact representativeness of state populations in our data depends on availability of MLO contact information. Collection of MLO contact information occurred from February through April of 2012. We used multiple styles of internet search including Google Maps, Google.com, Yellow Pages (YP.com) and Better Business Bureau (bbb.org) to locate MLO contact information.

Table 4. Mortgage Loan Originator Characteristics

	Number of Audits	Frequency	Overall Response Rate
Gender			
Female	1,916	36.98%	87.37%
Male	2,202	42.50%	84.92%
Not Identified	1,063	20.52%	80.53%
Race			
White	3,619	69.85%	86.57%
Non-White	273	5.27%	85.71%
Arabic	1	0.02%	100.00%
Asian	57	1.10%	80.70%
Black	90	1.74%	91.11%
Hispanic	115	2.22%	84.35%
Indian	7	0.14%	71.43%
Native American	3	0.06%	100.00%
Not Identified	1,289	24.88%	80.14%

Notes: We identified race and gender of MLOs by visually inspecting photographs when available on lender webpages. Our approach was conservative in identifying both gender and race: if we felt there was any room for argument about either, we categorized the demographic information as not identified. The not identified category includes all instances where there was ambiguity in assigning race and/or gender and when a photograph was not available. For the summary statistics shown, the Non-White race category includes Arabic, Asian, Black, Hispanic, Indian, and Native American, the breakdown of each is shown within the non-white category.

Table 5. Response Rate and Mortgage Loan Originator (MLO) Level Response to Race Differences

	Overall response rate			Response at MLO level				
	(1) White	(2) African American	(3) (1)-(2)	(4) Respond to neither	(5) Respond to both	(6) White only	(7) African American only	(8) (6)-(7)
All audits	68.31% [3540]	65.68% [3402]	2.63% $p=0.0022^{***}$	16.28% [632]	49.77% [1932]	17.88% [694]	16.07% [624]	1.81% $p=0.0573^*$
White MLOs	70.67% [2554]	67.41% [2443]	3.26% $p=0.0014^{***}$	14.49% [392]	51.96% [1406]	17.85% [483]	15.71% [425]	2.14% $p=0.0585^*$
Non-White MLOs	68.50% [187]	65.57% [179]	2.93% $p=0.2337$	13.04% [27]	49.28% [102]	20.77% [43]	16.91% [35]	3.86% $p=0.4282$
Missing Race MLOs	61.70% [799]	60.80% [780]	0.90% $p=0.3189$	21.98% [213]	43.76% [424]	17.34% [168]	16.92% [164]	0.42% $p=0.8693$
Male MLOs	69.16% [1496]	66.00% [1479]	3.16% $p=0.0124^{**}$	16.66% [277]	50.75% [844]	17.74% [295]	14.85% [247]	2.89% $p=0.0434^{**}$
Female MLOs	70.80% [1382]	67.23% [1264]	3.57% $p=0.0085^{***}$	12.98% [184]	51.48% [730]	18.83% [267]	16.71% [237]	2.12% $p=0.1964$
Missing Gender MLOs	62.04% [662]	62.23% [659]	-0.19% $p=0.5351$	21.35% [171]	44.69% [358]	16.48% [132]	17.48% [140]	-1.00% $p=0.6713$

Notes: The p-value represented in column (3) is from a one-sided t-test (alternative hypothesis of a positive difference) with a null hypothesis that the difference in average response rate is zero. The p-value reported in column (8) is from McNemar paired difference in proportions test. This test is designed for testing the difference in proportion of respondents for paired subjects, the test statistic is $\chi^2 = (N_{(Only W)} - N_{(Only AA)})^2 / (N_{(Only W)} + N_{(Only AA)})$, where N represents the number of MLOs only responding to one group. The test statistic has a chi-squared distribution, and we calculate all p-values accordingly. Number of MLOs shown in [], * = 0.10 significance, ** = 0.05 significance, *** = 0.01 significance.

Table 6. Response Rate and Mortgage Loan Originator (MLO) Level Response to Credit Score Differences

	(1)	(2)	(3)	(4)	(5)	(6)
Panel A: Response Rate Differences						
	High Credit	Low Credit	No Credit	(1)-(2)	(1)-(3)	(2)-(3)
Response Rate (includes all audits)	69.46% [2397]	65.76% [2276]	65.77% [2269]	3.70% <i>p</i> =0.0005***	3.69% <i>p</i> =0.0005***	-0.01% <i>p</i> =0.5024
Panel B: MLO Level Differences						
	Respond to neither	Respond to both	Respond to Higher only	Respond to Low/No only	(3)-(4)	
High vs. Low Credit	15.21% [132]	49.88% [433]	19.47% [169]	15.44% [134]	4.03% <i>p</i> =0.0506*	
High vs. No Credit	15.17% [129]	53.00% [450]	20.14% [171]	11.66% [99]	8.48% <i>p</i> =0.0000***	
Low vs. No Credit	17.80% [155]	50.17% [437]	17.22% [150]	14.81% [129]	2.41% <i>p</i> =0.2311	

Notes: The *p*-value represented in columns (4), (5), and (6) is from a one-sided *t*-test (alternative hypothesis of a positive difference) with a null hypothesis that the difference in average response rate is zero. The *p*-value reported in column (5) of Panel B is from McNemar paired difference in proportions test. This test is designed for testing the difference in proportion of respondents for paired subjects. The test statistic is $\chi^2 = (N_{(Only W)} - N_{(Only AA)})^2 / (N_{(Only W)} + N_{(Only AA)})$, where *N* represents the number of MLOs only responding to one group. The test statistic has a chi-squared distribution, and we calculate all *p*-values accordingly. Number of MLOs shown in [], * = 0.10 significance, ** = 0.05 significance, *** = 0.01 significance.

Table 7. Response Rate and Mortgage Loan Originator (MLO) Level Response by Race and Credit of Credit-Seeker

	Overall response rate			Response at MLO level				
	(1) Group 1	(2) Group 2	(3) (1)-(2)	(4) Respond to neither	(5) Respond to both	(6) White only	(7) African American only	(8) (6)-(7)
<u>Equal Credit Within Audit</u>								
White / African American	67.34% [1746]	65.72% [1704]	1.62% <i>p=0.1083</i>	16.31% [423]	49.36% [1280]	17.97% [466]	16.35% [424]	1.62% <i>p=0.1693</i>
<u>White Higher Credit Within Audit</u>								
White / African American	71.26% [300]	61.05% [257]	10.21% <i>p=0.0009***</i>	15.91% [67]	48.22% [203]	23.04% [97]	12.83% [54]	10.21% <i>p=0.0006***</i>
<u>African American Higher Credit Within Audit</u>								
White / African American	65.26% [248]	72.37% [275]	-7.11% <i>p=0.0173**</i>	15.00% [57]	52.63% [200]	12.63% [48]	19.74% [75]	-7.11% <i>p=0.0187**</i>

Notes: Equal Credit Within Audit implies credit category is the same (high, low, no) for a given audit. Higher credit score comparisons exclude low vs. no credit audits. Including no vs. low credit score audits makes differences between race even larger than the differences shown here. The p-value represented in column (3) is from a one-sided t-test (alternative hypothesis of a positive (or negative for African American Higher Credit) difference) with a null hypothesis that the difference in average response rate is zero. The p-value reported in column (8) is from McNemar paired difference in proportions test. This test is designed for testing the difference in proportion of respondents for paired subjects. The test statistic is $\chi^2 = \frac{(N_{(Only W)} - N_{(Only AA)})^2}{(N_{(Only W)} + N_{(Only AA)})}$, where N represents the number of MLOs only responding to one group. The test statistic has a chi-squared distribution, and we calculate all p-values accordingly. Number of MLOs shown in [], * = 0.10 significance, ** = 0.05 significance, *** = 0.01 significance.

Table 8. Intensity of Mortgage Loan Originator(MLO) Response

	(1) White	(2) African American	(3) (1)-(2)
<u>Panel A: Time Elapsed until Response</u>			
Time until response (h:mm)	8:20 (29:50)	9:23 (34:04)	1:03 $p=0.0879^*$
<u>Panel B: Length of Response</u>			
Character count, all audits	426.00 (356.35)	431.84 (388.34)	2.52 $p=0.3890$
Character count, dropping audits with equal (within 10 characters) length replies	448.74 (347.55)	431.84 (361.01)	16.90 $p=0.0986^*$
<u>Panel C: Follow-up Response</u>			
Follow-up e-mail received	6.91% [358]	5.17% [268]	1.74% $p=0.0001^{***}$
Number of follow-up e-mails received	2.19 (0.51)	2.25 (0.71)	-0.06 $p=0.9014$

Notes: Row 1 shows the average time elapsed between when an inquiry is sent and when a MLO reply is received, reported in h:mm format, these averages do not include e-mails where no reply was made. Rows 2 and 3 examine the character count for MLO responses. Row 2 examines all audits and includes counting non-response as zero characters, Row 3 excludes replies that were of equal length and does not count non-responses. Row 4 and 5 examine additional genuine e-mail responses after the first genuine e-mail response. Row 4 shows the percentage of MLOs who sent a follow-up e-mail, Row 5 shows the average number of follow-up e-mails received. In all cases, p-values are from standard difference in means z tests. Standard Deviations are reported in (), number of MLOs shown in [], *=0.10 significance, **= 0.05 significance, ***= 0.01 significance.

Table 9. Side-by-Side Comparison of Mortgage Loan Originator(MLO) Response Content

	(1) Neutral	(2) Prefer White	(3) Prefer African American	(4) (2)-(3)
Panel A: Author Blind Review				
MLOs Responding to Both Races (1,932 Matched Pairs)	57.25% [1106]	22.67% [438]	20.03% [387]	2.64% $p=0.0817^*$
Reason for Preference				
More Favorable Terms		5.94%	4.13%	$p=0.1200$
Friendliness		33.56%	32.82%	$p=0.4103$
Included More Details		46.58%	40.83%	$p=0.0484^{**}$
Explained the Process		5.48%	7.24%	$p=0.8498$
Un-preferred E-mail was Negative		5.25%	8.01%	$p=0.9451$
Facilitated the Transaction		27.85%	30.75%	$p=0.8193$
Un-preferred E-mail Steered or was Pushy		1.37%	0.78%	$p=0.2059$
Other		9.82%	10.85%	$p=0.6873$
Panel B: Outside Reviewer Blind Review				
MLOs Responding to Both Races (1,932 Matched Pairs)	43.94% [849]	29.19% [564]	26.40% [510]	2.79% $p=0.1058$
Reason for Preference				
More Favorable Terms		5.67%	5.10%	$p=0.3384$
Friendliness		47.16%	45.29%	$p=0.2698$
Included More Details		54.26%	51.57%	$p=0.1892$
Explained the Process		20.04%	18.82%	$p=0.3082$
Un-preferred E-mail was Negative		7.98%	7.45%	$p=0.6268$
Facilitated the Transaction		26.24%	26.47%	$p=0.5340$
Un-preferred E-mail Steered or was Pushy		2.13%	1.37%	$p=0.1743$
Other		3.90%	5.29%	$p=0.8628$

Notes: Side-by-Side comparison uses the visual basic interface shown in Appendix 2 for all MLOs responding to both e-mails in matched pairs with clients of different race. Column (2) totals includes all instances where whites were preferred or strongly preferred, Column (3) includes all instances where African Americans were preferred or strongly preferred. All indications of preference in panel A are judged by the authors in a blind review where information about clients and MLOs is masked. All indications of preference in panel B are indicated by outside reviewers in a blind review where information about clients and MLOs is masked. The p-value represented in column (4) is from a one-sided t-test (alternative hypothesis of a positive difference) with a null hypothesis that the difference in average response rate is zero. Percentages for reasons indicated for preference do not sum to one because graders were allowed to indicate multiple reasons for preference. See Appendix 3 for instructions given to graders and definitions of reasons for preference. Number of MLOs shown in [], *=0.10 significance, **= 0.05 significance, ***= 0.01 significance.

Table 10: Name Robustness and National Popularity

Name	Response Rate	Different than own race	National Popularity
Jake Krueger	72.62%	p=0.0422	140
Brett Nelson	70.92%	p=0.2304	84
Brendan Nelson	70.09%	p=0.3988	155
Ethan Schmitt	69.70%	p=0.5146	58
Luke Krueger	69.65%	p=0.5215	118
Tyrone Washington	68.80%	p=0.1324	260
Jermaine Booker	68.01%	p=0.2948	306
Zachary Miller	67.98%	p= 0.8761	22
Reginald Jackson	67.96%	p=0.3082	230
Darnell Jackson	67.42%	p=0.4379	353
Conor Schmitt	67.13%	p=0.5834	458
Spencer Miller	66.26%	p=0.3510	105
Jamal Washington	66.08%	p=0.8533	247
Seth Baker	65.77%	p=0.2352	102
Kadeem Jefferson	64.71%	p=0.6751	538
Maxwell Baker	64.66%	p=0.0901	188
DaQuan Booker	63.54%	p=0.3184	709
Jerome Jefferson	63.05%	p=0.2101	264
Terrell Banks	62.93%	p=0.2110	282
DaShawn Banks	62.62%	p=0.1599	732

Notes: P-value is for a difference in means t-test between the response rate for each name and names of the same race. National popularity ranking comes from the Social Security Administration website using counts of baby names from 1990 at: <http://www.ssa.gov/cgi-bin/popularnames.cgi>

Table 11. Results excluding popular or unique names.

	Overall response rate			Response at MLO level				
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	White	African American	(1)-(2)	Respond to neither	Respond to both	White only	African American only	
Exclude Jake Krueger (5% level different than own race)	67.83% [3158]	65.68% [3402]	2.15% <i>p</i> =0.0119**	16.41% [571]	49.71% [1730]	17.53% [610]	16.35% [569]	1.18% <i>p</i> =0.2440
Exclude Jake Krueger and Maxwell Baker (10% different than own race)	68.46% [2835]	65.68% [3402]	2.78% <i>p</i> =0.0023***	15.87% [490]	50.31% [1553]	17.75% [548]	16.07% [496]	1.68% <i>p</i> =0.1144
Exclude DaShawn Banks and Jake Krueger (highest white, lowest AA)	67.83% [3158]	66.02% [3072]	1.81% <i>p</i> =0.0321**	16.27% [508]	49.98% [1561]	17.61% [550]	16.14% [504]	1.47% <i>p</i> =0.1657
Exclude Zachary Miller and DaShawn Banks (most popular white, least popular AA)	68.35% [3194]	66.02% [3072]	2.33% <i>p</i> =0.0083***	16.02% [503]	50.94% [1599]	17.43% [547]	15.61% [490]	1.82% <i>p</i> =0.0820*
Exclude Zachary Miller and Ethan Schmitt (2 most popular)	68.18% [2826]	65.68% [3402]	2.50% <i>p</i> =0.0054***	16.29% [503]	50.47% [1558]	17.27% [533]	15.97% [493]	1.30% <i>p</i> =0.2234
Exclude DaShawn Banks and DaQuan Booker (2 least popular)	68.31% [3540]	66.35% [2727]	1.96% <i>p</i> =0.0224**	15.62% [482]	50.36% [1554]	17.85% [551]	16.17% [499]	1.68% <i>p</i> =0.1155

Notes: The *p*-value represented in column (3) is from a one-sided *t*-test (alternative hypothesis of a positive difference) with a null hypothesis that the difference in average response rate is zero. The *p*-value reported in column (8) is from McNemar paired difference in proportions test. This test is designed for testing the difference in proportion of respondents for paired subjects, the test statistic is $\chi^2 = (N_{(Only\ W)} - N_{(Only\ AA)})^2 / (N_{(Only\ W)} + N_{(Only\ AA)})$, where *N* represents the number of MLOs only responding to one group. The test statistic has a chi-squared distribution, and we calculate all *p*-values accordingly. Number of MLOs shown in [], * = 0.10 significance, ** = 0.05 significance, *** = 0.01 significance.